

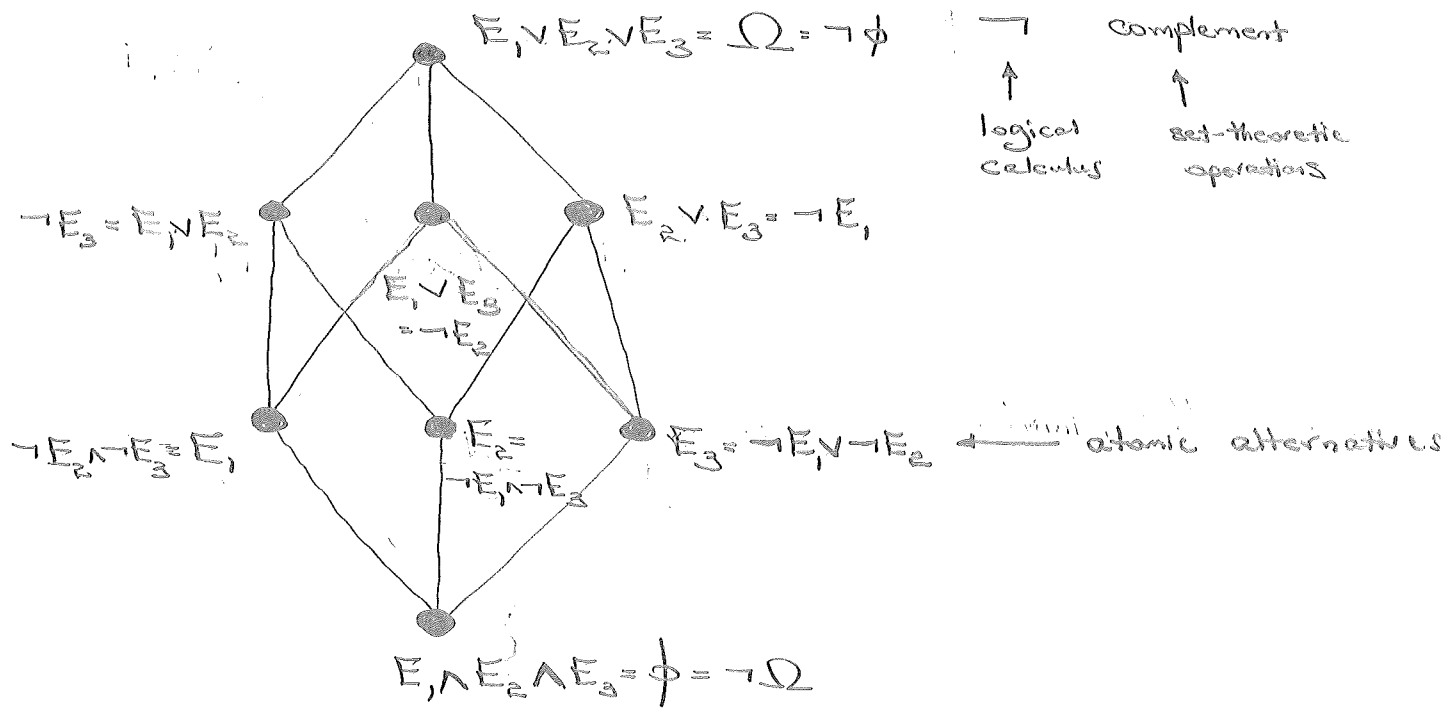
Quantum information theory

Lecture 1

Probabilities and laws of large numbers

Boolean lattice of events
propositions
alternatives

$\vee$ $\cup$
 $\wedge$ $\cap$
 $\neg$ complement
 $\uparrow$ logical calculus
 $\uparrow$ set-theoretic operations



Axioms of probabilities

Foundations: See Dutch-book documents on web page.

- ① $0 \leq p(E) \leq 1$ ← measure on the event space
- ② E is certain $\iff p(E) = 1$

③ Additivity for mutually exclusive events:

$$p(E_1 \vee E_2) = p(E_1) + p(E_2) \text{ if } E_1 \wedge E_2 = \phi$$

④ Bayes's rule: $p(E_1 \wedge E_2) = p(E_1 | E_2) p(E_2)$

Or normalization: $\begin{matrix} \nearrow \text{②+③} \\ \searrow \text{normalization + ②+④} \end{matrix} \Rightarrow p(E) + p(\neg E) = p(E \vee \neg E) = p(\Omega) = 1$

$$\Rightarrow E = E_1 \vee E_2$$

$$p(E) = \underbrace{[p(E_1 | E) + p(E_2 | E)]}_{1} p(E) = \underbrace{p(E | E_1)}_{1} p(E_1) + \underbrace{p(E | E_2)}_{1} p(E_2) = p(E_1) + p(E_2)$$

(E)

Random variable X has values $\underbrace{x_1, \dots, x_N}_{\text{atomic alternatives}}$

$$P_X(x)$$

Two random variables: X, Y

Joint probability: $P_{X,Y}(x,y)$

Marginal probabilities:)

$$P_X(x) = \sum_y P_{X,Y}(x,y)$$
$$P_Y(y) = \sum_x P_{X,Y}(x,y)$$

Conditional probabilities:

$$\underbrace{P_{X|Y}(x|y)}_{\substack{\text{probability of } x, \\ \text{given that } Y=y}} P_Y(y) = P_{X,Y}(x,y) = P_{Y|X}(y|x) P_X(x)$$

Bayes's rule

We almost always omit the subscripts that indicate which random variable(s) the probability applies to, relying on the argument to provide that information.

$$p \leftrightarrow P \leftrightarrow P_r$$

Laws of large numbers

① Chebyshev inequality. For any nonnegative random variable X with finite mean, $P(X > a) \leq \langle X \rangle / a$.

Proof:

$$P(X > a) = \int_a^{\infty} dx p(x) \leq \int_a^{\infty} dx \frac{x}{a} p(x) \leq \int_0^{\infty} dx \frac{x}{a} p(x) = \frac{\langle X \rangle}{a}$$

⇒ ② For any random variable X with finite mean and variance,

$$P(|X - \langle X \rangle| > a) = P((X - \langle X \rangle)^2 > a^2) \leq \frac{\langle (\Delta X)^2 \rangle}{a^2}$$

$$\left(\begin{array}{c} \text{variance} \\ \text{of } X \end{array} \right) = \langle (\Delta X)^2 \rangle = \langle (X - \langle X \rangle)^2 \rangle = \langle X^2 \rangle - \langle X \rangle^2 = \sigma_X^2$$

↑
alternate notation
we won't use
much

Now let $X_1, \dots, X_N$ be independent, identically distributed (i.i.d.) random variables with finite mean $\langle x \rangle$ and variance $\langle (\Delta x)^2 \rangle$. The sample mean

$$S = \frac{1}{N} \sum_{\ell=1}^N X_{\ell}$$

has mean

$$\langle S \rangle = \frac{1}{N} \sum_{\ell=1}^N \langle X_{\ell} \rangle = \langle x \rangle$$

↓
= $\langle x \rangle$

and second moment

$$\begin{aligned}\langle S^2 \rangle &= \frac{1}{N^2} \sum_{\ell, m} \langle x_\ell x_m \rangle = \langle x \rangle^2 + \frac{1}{N} \langle (\Delta x)^2 \rangle \\ &= \langle (\langle x \rangle + \Delta x_\ell)(\langle x \rangle + \Delta x_m) \rangle \\ &= \langle x \rangle^2 + \underbrace{\langle \Delta x_\ell \Delta x_m \rangle}_{= \delta_{\ell m} \langle (\Delta x)^2 \rangle \leftarrow \text{i.i.d.}} \\ &= \langle x \rangle^2 + \sum_{\ell m} \langle (\Delta x)^2 \rangle\end{aligned}$$

Thus the variance of the sample mean is

$$\langle (\Delta S)^2 \rangle = \langle S^2 \rangle - \langle S \rangle^2 = \frac{1}{N} \langle (\Delta x)^2 \rangle$$

Weak law of large numbers I.

③

Weak law of large numbers II.

$$P(|S - \langle x \rangle| > \epsilon) \leq \frac{\langle (\Delta S)^2 \rangle}{\epsilon^2} = \frac{\langle (\Delta x)^2 \rangle}{N \epsilon^2}$$

↑
Chebyshev inequality

④

This weak law says that

$$\lim_{N \rightarrow \infty} P(|S_N - \langle x \rangle| \leq \epsilon) = 1 \text{ for any } \epsilon > 0.$$

To prove the asymptotic equipartition property (AEP),

we will use this form of the weak law with a random variable that takes on values

$$p_\ell = -\log p(x_\ell).$$

⑤ Relation to central-limit theorem?

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Now let X be a discrete random variable that can take on D values $x_1, \dots, x_D$. Let $p_j = p(x_j)$ be the probability of the j th value. In N trials, the probability of a sequence $x_1, \dots, x_N$ is

$$p(x_1, \dots, x_N) = p(x_1) \dots p(x_N).$$

Let n_j be the number of occurrences of j ($\sum_{j=1}^D n_j = N$). The probability of occurrence numbers $n_1, \dots, n_D$ is

$$p(n_1, \dots, n_D) = \frac{N!}{n_1! \dots n_D!} p_1^{n_1} \dots p_D^{n_D}$$

(i) Normalization: $\sum_{n_1, \dots, n_D} p(n_1, \dots, n_D) = (p_1 + \dots + p_D)^N = 1$

(ii) Means:

$$\begin{aligned} \langle n_j \rangle &= \sum_{n_1, \dots, n_D} n_j p(n_1, \dots, n_D) = p_j \frac{\partial}{\partial p_j} \left(\sum_{n_1, \dots, n_D} p(n_1, \dots, n_D) \right) = N p_j \\ &= p_j \frac{\partial}{\partial p_j} (p_1 + \dots + p_D)^N \\ &= N p_j (p_1 + \dots + p_D)^{N-1} \\ &= N p_j \end{aligned}$$

(iii) Correlation matrix:

$$\langle \eta_j \eta_k \rangle = \sum_{\eta_1, \dots, \eta_D} \eta_j \eta_k p(\eta_1, \dots, \eta_D) = \underbrace{p_j \frac{\partial}{\partial p_j} \left(p_k \frac{\partial}{\partial p_k} \left(\sum_{\eta_1, \dots, \eta_D} p(\eta_1, \dots, \eta_D) \right) \right)}_{= N p_k (p_1 + \dots + p_D)^{N-1}}$$

↑
end-moment matrix

$$\begin{aligned} \langle \eta_j \eta_k \rangle &= N^2 p_j p_k + N (p_j \delta_{jk} - p_j p_k) \\ &= N p_j \delta_{jk} \binom{N-1}{\quad} + N(N-1) p_j p_k \binom{N-2}{\quad} \\ &= N^2 p_j p_k + N (p_j \delta_{jk} - p_j p_k) \end{aligned}$$

The correlation matrix is

$$\begin{aligned} \langle \Delta \eta_j \Delta \eta_k \rangle &= \langle (\eta_j - \langle \eta_j \rangle) (\eta_k - \langle \eta_k \rangle) \rangle \\ &= \langle \eta_j \eta_k \rangle - \langle \eta_j \rangle \langle \eta_k \rangle \\ &= N (p_j \delta_{jk} - p_j p_k) \end{aligned}$$

Notice that for $j \neq k$,

$$\langle \Delta \eta_j \Delta \eta_k \rangle = - N p_j p_k$$

↑
anticorrelation
because occurrence
numbers must add
to N.

The variances are

$$\langle (\Delta \eta_j)^2 \rangle = N p_j (1 - p_j)$$

We want to work with the frequency of occurrence of each alternative:

$$f_j = \frac{\eta_j}{N}$$

$$\langle f_j \rangle = \frac{\langle \eta_j \rangle}{N} = p_j$$

$$\langle f_j f_k \rangle = \frac{\langle \eta_j \eta_k \rangle}{N^2} = p_j p_k + \frac{1}{N} (p_j \delta_{jk} - p_j p_k)$$

$$\langle \Delta f_j \Delta f_k \rangle = \frac{1}{N} (p_j \delta_{jk} - p_j p_k)$$

$$\langle (\Delta f_j)^2 \rangle = \frac{p_j (1 - p_j)}{N}$$

④ Weak law of large numbers. III.

$$P(|f_j - p_j| \leq \epsilon, j=1, \dots, D) \geq 1 - \frac{1 - \sum_{j=1}^D p_j^2}{N\epsilon^2}$$

probability that all the frequencies lie within ϵ of the corresponding probability

We could do the frequency of each alternative separately but here we ask a question about all frequencies.

This weak law says that

$$\lim_{N \rightarrow \infty} P(|f_j - p_j| \leq \epsilon, j=1, \dots, D) = 1 \text{ for any } \epsilon > 0.$$

Proof:

Sphere of radius ϵ inside the hypercube

$$\begin{aligned} P(|f_j - p_j| \leq \epsilon, j=1, \dots, D) &\geq P\left(\sum_j (f_j - p_j)^2 \leq \epsilon^2\right) \\ &\quad \text{hypercube with sides of length } 2\epsilon \\ &= 1 - P\left(\sum_j (f_j - p_j)^2 > \epsilon^2\right) \\ &\leq \frac{\sum_j \langle (\Delta f_j)^2 \rangle}{\epsilon^2} = \frac{1 - \sum p_j^2}{N\epsilon^2} \\ &\quad \uparrow \text{Chebyshev inequality} \\ &\geq 1 - \frac{1 - \sum_{j=1}^D p_j^2}{N\epsilon^2} \end{aligned}$$

The weak laws refer to a limit of a sequence of probabilities. The strong law of large numbers deals with limits of the frequencies themselves. To illustrate this, let's consider the case of binary alternatives, 0

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and 1. A sequence $x_1, \dots, x_N$, has a frequency of occurrence of 1 given by

$$f_N = \frac{1}{N} \sum_{n=1}^N x_n.$$

The strong law deals directly with the limit of this frequency sequence and says that

$$P\left(\lim_{N \rightarrow \infty} f_N = p\right) = 1.$$

↑
probability for 1

Many sequences don't have a frequency limit at all, so you have to show that there are sequences that do have a frequency limit and that these have all the probability (this latter task requires formulating a probability measure directly on the space of infinite sequences).

We never seem to use the strong law because ultimately we're interested in limits of probabilities on finite sets, which we can understand, not probabilities of limits, which have only a formal significance.